Effects of Nondeterminism on the Predicted Speedup of Scheduling of Low Level Computer Vision Algorithms on Networks of Heterogeneous Machines

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Abstract

Defining an optimal schedule for arbitrary algorithms on a network of heterogeneous machines is an NP complete problem. This paper focuses on data parallel deterministic neighborhood computer vision algorithms. This focus enables the linear time definition of a schedule which minimizes the distributed execution time by overlapping computation and communication cycles on the network. The static scheduling model allows for any speed machine to participate in the concurrent computation but makes the assumption of a master/slave control mechanism using a linear communication network. We investigate the limitations of the static scheduling model based on statistical descriptions of the model parameters. Using statistical models, an approximation of the schedule length density function is derived. This statistical model is used to establish better approximations of schedule length.

Keywords: Computer Vision, Heterogeneous Scheduling, Distributed Algorithms, Nondeterminism

1. Introduction

In the past, many studies have been performed analyzing the capabilities of various parallel processor – vision algorithm mappings. Thorough surveys can be found in [17][3][2][4][5][16]. Most of these efforts focus on the mapping of a single machine to a single algorithm, or mapping a suite of algorithms to a single architecture [17]. Most of the conclusions made in these studies are based on the architectural similarities between the hardware communication configurations and the communication patterns inherent in the vision algorithm (i.e. vision tasks tend to have highly regular communication). Recent research efforts have discussed the mapping of computer vision tasks to networks of workstations with the assumption of homogeneous workstation clusters [11][7]. Additional efforts have focused on scheduling suites of independent programs onto networks of heterogeneous machines [10]. This paper relaxes the assumption of homogeneous workstation clusters and independent program suites. It focuses on the distribution of a single program on a set of architectures connected by the PVM message passing library [6][13]. A framework is presented for the polynomial time static scheduling of deterministic local communication algorithms onto a suite of heterogeneous machines using linear communication. In order to evaluate the effectiveness of the static schedule, a nondeterministic model is defined.

The paper is organized as follows. Section 2 presents the background and motivation. Section 3 develops the analytical description of the scheduling process and introduces a set of conditions necessary for the minimization of the execution time. Section 4 presents the scheduling algorithm. Section 5 presents several low level computer vision tasks and their corresponding scheduling models. Section 6 introduces a statistical model of the schedule. Section 7 demonstrates the use of the statistical model for predicting schedule length of a distributed convolution algorithm. Section 8 presents conclusions and future work.

2. Background

Efforts have been made to develop an Automatic Visual Inspection System (AVIS) for use with various aerospace engine components. These components include high pressure turbine blades (aircraft), injector baffles, oxidation posts and annulus rings (spacecraft). AVIS utilizes several scales of information abstracted from the original image with each scale requiring a set of low level vision operations [8][14]. The realization of the AVIS paradigm is limited by the tremendous computational burden of these low level vision operations. These algorithms include convolution, difference of Gaussian filtering, morphological filtering, Fourier transform, and Hough transform. In order to increase the speed of AVIS, distributed solutions were investigated. Defining an effective distribution onto the various machines available on the LAN requires models of algorithm decomposition, communication mechanisms, and machine speed for a given algorithm.

3. Schedule Model

Several assumptions are made in the definition of the schedule model. The communication time, execution time, and algorithm decomposition are assumed to be linear. The control structure is assumed to be master/slave with no interslave communication. Based on these assumptions, we construct:

$$\mathfrak{I}(n)$$
 – a scheduling of N homogeneous tasks onto n machines, Eq.3.1

where
$$N = \sum_{i=1}^{i=n} \eta_i$$
 , with each slave machine(i) receiving a task size of η_i .

The time to send data to the slave(i) in the schedule is described by:

$$T_{s(i)} = T_{s(i)}^{0}(\text{overhead}) + T_{s(i)}^{\eta_{i}}(\text{transmit}).$$
 Eq.3.2

Similarly, the time to receive data from slave(i) is described by:

$$T_{r(i)} = T_{r(i)}^{0}(\text{overhead}) + T_{r(i)}^{\eta_{i}}(\text{transmit})$$
. Eq.3.3

The computation time for slave(i) can also be described in a similar manner:

$$T_{c(i)} = T_{c(i)}^{\eta_i}(\text{compute } \eta_i)$$
. Eq.3.4

These parameters can be defined in terms of communication bandwidth β_i (secs/pixel), communication overhead, α_i (secs), and execution rate, γ_i (secs/operation) for each slave machine used in the schedule. Because the model is linear, the computation and communication times are proportional to their corresponding neighborhoods (data sizes), as described below.

$$T^1_{s(i)} = \phi_s \times \beta_i$$
, Eq.3.5

where ϕ_s is the neighborhood associated with sending a task size 1 to a slave(i).

$$T^{0}_{s(i)} = \phi^{0}_{s} \times \beta_{i} + \alpha_{i},$$
 Eq.3.6

where ϕ_s^0 is the send overhead neighborhood.

$$T^{1}_{r(i)} = \phi_r \times \beta_i$$
, Eq.3.7

where ϕ_r is the neighborhood associated with task size 1 received from slave(i) .

$$T^0_{\ r(i)} = \ \varphi^0_{\ r} \times \beta_i + \alpha_i, \quad \text{where } \varphi^0_r \text{ is the receive overhead neighborhood} \ .$$
 Eq.3.8

$$T^{1}_{c(i)} = \phi_{c} \times \gamma_{i}$$
, Eq.3.9

where ϕ_c is the neighborhood associated with computing a single task on slave(i) .

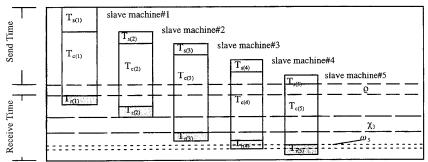


Fig.1 Depiction of Master/Slave Communication and Execution Model

As seen in Fig.1, the linear model enables a direct analysis of the optimality of a given schedule. The degree of nonconcurrency in the schedule is indicated by the values of ϱ , ω , and χ . The parameter ϱ is a measure of the time spent by the master machine waiting for the first processed data partition to be returned,

$$\varrho = T_{c(1)} - \sum_{i=2}^{n} T_{s(i)}$$
 Eq.3.10

The parameter ω_i is a measure of slave wait time caused by contention with the previous slave machine for the linear communication channel,

$$\omega_{i} \ = \ \left\lfloor \ T_{c(i-1)} + T_{r(i-1)} - \left(\ T_{s(i)} + T_{c(i)} \ \right) \ \right\rfloor \ , \ \ \text{where} \ \left\lfloor f \right\rfloor \ = \ \text{max} \ (0,f) \ . \quad \text{Eq.3.11}$$

The parameter χ_i is a measure of the master wait time caused by excess computation of machine(i) after machine(i-1) has completed returning its data partition. The effects on the schedule time can be reduced by previous ω_i and ϱ values, i.e.

$$\chi_{i} = [T_{s(i)} + T_{c(i)} - T_{c(i-1)} - T_{r(i-1)} - \lambda_{i}].$$
 Eq.3.12

The λ_i term represents the data queue waiting to be returned from the slave processes due to previous contention for the communication channel,

$$\lambda_{i} = \lfloor \lambda_{i-1} + \omega_{i-1} - \chi_{i-1} \rfloor, \ \lambda_{1} = \lfloor -\varrho \rfloor.$$
 Eq.3.13

As one would expect, large values of ϱ , ω and χ result in suboptimal schedules. Typically, these values cannot be eliminated due to the constraint of integer task size. The total execution time for a given schedule is:

$$T(n) = \sum_{i=1}^{n} T_{s(i)} + \sum_{i=1}^{n} T_{r(i)} + \lfloor \varrho \rfloor + \sum_{i=2}^{n} \chi_{i}.$$
 Eq.3.14

In order to minimize this execution time, a set of criteria must be established to explicitly define the effects of the values ϱ, ω_i , and χ_i . In a previous paper [18], a set of conditions were derived which define these criteria. Based on these conditions, a linear time static scheduling algorithm was presented.

4. Scheduling Algorithm

As developed in [18], the minimum schedule length (Eq.3.14) for n machines can be found in which $\omega_i=0$ and $\chi_i=0$. This solution corresponds to $T_{c(i)}+T_{r(i)}=T_{c(i+1)}+T_{s(i+1)}$. Using

this relation in conjunction with the definition of total task size, $\sum_{i=1}^{i=n} \eta_i = N$, the n values of η_i can be determined explicitly using Gaussian elimination on the n linear equations shown in

 η_i can be determined explicitly using Gaussian elimination on the n linear equations shown in Fig.2.

(1)
$$T_{c (1)} + T_{r(1)} = T_{c(2)} + T_{s(2)}$$

 \vdots
 $(n-1) T_{c (n-1)} + T_{r(n-1)} = T_{c(n)} + T_{s(n)}$
(n) $\sum_{i=1}^{n} \eta_i = N$ Eq.4.1

Fig.2 Set of linear equations defining schedule with $\chi_i=0,\,\omega_i=0$ for all i

Although no closed form solution for n exists, an appropriate number of slave machines can be determined by the evaluation of ϱ as described by conditions developed in [18]. Because the solution to Eq.4.1 is based on the deterministic values of γ_i , α_i and β_i , the static schedule is susceptible to any observed deviation of these parameters. An analysis of these nondeterministic effects will be presented in section 6.

5. Low Level Vision Modeling

The low level computer vision tasks needed for the AVIS computations are convolution, difference of Gaussian, Fourier transform, Hough transform, and morphological filtering operations. Although explication of these algorithms is beyond the scope of this paper, thorough discussions can be found in [1][4][5][12][15][16][17]. For the purpose of this study, the key algorithmic elements are contained in the model parameters, $\gamma_i, \varphi_c, \varphi_s, \varphi_c, \varphi_s^0$. These algorithm specific parameters are presented in Fig.3.

	Convolution	DoG	Morphological	2D FFT	Hough
Υį	secs mult + add	secs mult + add	secs 2logical Ops	secs cmplx(mult + add)	secs mult + trig_lookup
Фс	M ² N	$(M_1^2 + M_2^2)N$	BN	Nlog ₂ N	#image pixels
фѕ	N	N	N	N	0
фr	N	N	N	N	1
ф	$(M^2 + (M-1) N)$	$(M_1^2 + M_2^2 + (M_1 - 1) N)$	(B+ (M-1)N)	0	N ²
ϕ_r^0	0	0	0	0	0

where Mi is the size of the square kernel, N is the image size and B is the maximum size of the rectangular kernel.

Fig.3 Table of scheduling parameters for the low level vision algorithms

6. Statistical Properties

In order to proceed with a statistical analysis, it is necessary to make assumptions about the density distribution of the schedule parameters α_i , β_i and γ_i . Normal densities are assumed, and the sample means and variances of these parameters are used as unbiased estimates of the mean and variances. These means and variances are defined as:

 μ_{α} = sample mean of communication overhead to slave machine i

 σ_{α_i} = standard deviation of communication overhead to slave machine i

 $\mu_{\beta_{i}}$ = sample mean of communication bandwidth to/from slave machine is

 σ_{β} = standard deviation of communication bandwidth to/from slave machine i

= sample mean of execution rate of slave machine i

 $\sigma_{\gamma_{\rm c}}$ = standard deviation of execution rate of slave machine i

We represent the communication overhead as a random variable O_i , the communication bandwidth as random variable C_i, and the execution rate as random variable G_i. The density distributions for these random variables are given in Eq.6.1.

$$f(O_i) = \frac{1}{\sigma_{\alpha_i}\sqrt{2\pi}} \ e^{-\frac{\left(x-\mu_{\alpha_i}\right)^2}{2\sigma_{\alpha_i}^2}}, \quad f(C_i) = \frac{1}{\sigma_{\beta_i}\sqrt{2\pi}} \ e^{-\frac{\left(x-\mu_{\beta_i}\right)^2}{2\sigma_{\beta_i}^2}}, \\ f(G_i) = \frac{1}{\sigma_{\gamma_i}\sqrt{2\pi}} \ e^{-\frac{\left(x-\mu_{\gamma_i}\right)^2}{2\sigma_{\gamma_i}^2}} \ Eq. 6.1$$

These variables are assumed to be normal and independent, hence any linear combination of them is also a normal random variable with variance and mean defined as:

$$\sigma_{\sum_{i=1}^{n}c_{i}x_{i}}^{2} = variance\Biggl[\sum_{j=1}^{n}c_{i}x_{i}\Biggr] = \sum_{i=1}^{n}c_{i}^{2}\sigma_{i}^{2} \ , \\ \mu_{\sum_{i=1}^{n}c_{i}x_{i}} = mean\Biggl[\sum_{i=1}^{n}c_{i}x_{i}\Biggr] \ = \sum_{i=1}^{n}c_{i}\mu_{i} \ , \\ Eq. 6.2$$

where x; is a normal independent random variable and c; is a linear multiplier.

Representing Eq. 3.14 in terms of these random variables is complicated by the parameter λ_i (as defined in Eq.3.13). Assuming that $\lambda_i = 0$, an upper approximation of T(n) can be defined:

$$T(n) \approx \sum_{i=1}^{n} T_{s(i)} + \sum_{i=1}^{n} T_{r(i)} + \lfloor \varrho \rfloor + \sum_{i=2}^{n} \lfloor T_{s(i)} + T_{c(i)} - T_{c(i-1)} - T_{r(i-1)} \rfloor \cdot \text{Eq.6.3}$$
Representing the observed time, $T(n)$, as a random variable Z , we examine the upper approximation of Z in terms of the readom variables $W(X)$ and Y .

tion of Z in terms of the random variables W, X, and Yi.

$$Z \approx W + \lfloor X \rfloor + \sum_{i=2}^{n} \lfloor Y_i \rfloor$$
. Eq.6.4

These random variables are defined as combinations of the random variables representing the startup overhead, communication bandwidth, and execution rates of the slave machines, O_i, C_i, and Gi, respectively. The definition for W and its corresponding density is given in Eq.6.5. The definition for X and its corresponding density is given in Eq.6.6, and the definition of Y₁ and its corresponding density is given in Eq.6.7.

$$W = \sum_{i=1}^{n} \left(O_i + (\eta_i \phi_s + \phi_s^0) C_i \right) + \sum_{i=1}^{n} \left(O_i + (\eta_i \phi_r + \phi_r^0) C_i \right)$$
 Eq.6.5

$$f(W) = \frac{1}{\sigma_W \sqrt{2\pi}} \ e^{-\frac{\left(x - \mu_W\right)^2}{2\sigma_W^2}} \ , \ \text{where} \ \mu_W, \ \sigma_W \ \text{are given by Eq.6.2} \ .$$

$$X = (\eta_1 \ \phi_c) \ G_1 - \sum_{i=2}^{n} (O_i + (\eta_i \ \phi_s + \phi_s^0) \ C_i)$$
 Eq.6.6

$$f(X) = \frac{1}{\sigma_X \sqrt{2\pi}} \; e^{-\frac{\left(x-\mu_X\right)^2}{2\sigma_X^2}} \; , \; \text{where} \; \mu_X \; , \; \sigma_X \; \text{ are given by Eq.6.2} \; .$$

$$\begin{split} Y_i &= \; (\eta_i \varphi_c) G_i - (\eta_{i-1} \; \varphi_c) G_{i-1} + \left(O_i + (\eta_i \varphi_s \; + \; \varphi_s^0) C_i \right) - \; \left(O_{i-1} + (\eta_{i-1} \varphi_r + \; \varphi_r^0) \; C_i \right), \\ f(Y_i) &= \; \frac{1}{\sigma_{Y_i} \sqrt{2\pi}} \; e^{-\frac{\left(x - \mu_{Y_i}\right)^2}{2\sigma_{Y_i}^2}} \; , \\ &\qquad \qquad \qquad \text{Eq.6.7} \end{split}$$

As described in Eq.6.5, Eq.6.6, and Eq.6.7, the density functions for W, X, and Y are normal. The mean for W is strictly greater than zero, whereas the means for X and Y_i can be greater or less than zero. If the X and Y_i random variables were not truncated at zero as indicated in Eq.6.4, the composite density function for the upper bound of Z would be a simple summation of normal variables. The truncated normal density is defined with a delta function at the origin as described in Eq.6.8.

$$\begin{split} \lfloor X \rfloor \; &= \; \begin{cases} X \quad \text{if } X \geq 0 \\ 0 \quad \text{if } X < 0 \end{cases}, \quad f(\lfloor X \rfloor) \; = \; \begin{cases} \theta_X \, \delta(x) \, + \, \frac{1}{\sigma_X \sqrt{2\pi}} \, \, e^{-\frac{\left(x - \mu_X\right)^2}{2\sigma_X^2}} \, \, \text{if } x \geq 0 \quad \text{Eq.6.8} \\ 0 \qquad \qquad \text{if } x < 0 \end{cases} \\ \text{where } \; \theta_X = \; \int\limits_{-\infty}^0 \, \frac{1}{\sigma_X \sqrt{2\pi}} \, \, e^{-\frac{\left(x - \mu_X\right)^2}{2\sigma_X^2}} \, dx, \text{and } \delta(x) \text{ is the Dirac delta function }. \end{split}$$

The inclusion of the truncated normal curve into the composite density function makes the definition of f(Z) more difficult. Assuming the three random variables, W, [X], $[Y_i]$ are in-

dependent, we can define the density of their sum as the convolution of their densities [19]. As such, the density of the upper bound of Z can be defined as:

$$\begin{split} f_z(Z) &= \left(\ f(W) * \ f(\lfloor X \rfloor) \right) * f(\lfloor Y_2 \rfloor) * f(\lfloor Y_3 \rfloor) \ ... * f(\lfloor Y_n \rfloor) \right), \\ \text{where } f(W) * f(\lfloor X \rfloor) &= \int\limits_{-\infty}^{\infty} f_W(Z - X) f_{\lfloor X \rfloor}(X) dX \ . \end{split}$$

The difficulty introduced by the truncated normal curve is based on the integral of Eq.6.9 . The limits of integration are redefined as:

$$f(W) * f([X]) = \int_{0}^{\infty} f_{W}(Z - X) \left[\theta_{X} \delta(X) + \frac{1}{\sigma_{X} \sqrt{2\pi}} e^{-\frac{(X - \mu_{X})^{2}}{2\sigma_{X}^{2}}} \right] dX. \quad \text{Eq.6.10}$$

This integral has no closed form solution. Numerical solutions, however, offer a simple solution, and can quickly define a numerical density for the lower bound of random variable Z. After this density is defined, numerical integration can be performed on f(Z) to define confidence intervals on the predicted schedule time.

7. Experiments

As explained in section 6, an approximate density function for the schedule time can be defined using Eq.6.9. Given a schedule derived using Eq.4.1, the probability that the schedule will exhibit an observed speedup can be obtained. To demonstrate the use of this information two disjoint slave sets were arbitrarily defined. These two slave sets were used to examine the correlation between the observed schedule times and those predicted by the deterministic and nondeterministic models.

The schedule length density approximations for slave set 1 and slave set 2 are depicted below in Fig. 4. In addition to the density, the observed schedule time (averaged over 30 trials), the predicted deterministic model time and the density means are presented. These values are summarized below in Fig. 5 for several kernel sizes of the convolution algorithm. In general the observed schedule times were greater than that predicted by the deterministic model and less than the nondeterministic mean. An exception to this observation can be found in the 7x7 convolution for slave set 2 in Fig. 5.

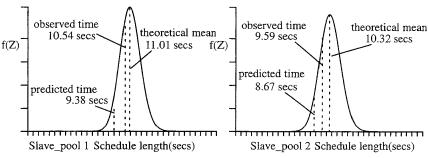


Fig.4 Density approximations of two disjoint slave pools for 17x17 convolution

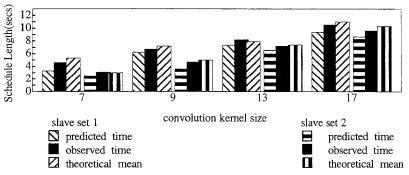


Fig.5 Predicted time, observed time and theoretical mean for slave sets 1&2

8. Conclusions

Focusing on linearly partitionable problems enables the definition of optimal scheduling conditions. These conditions enable a schedule to be defined in linear time. These conditions, however, are based on the assumption of deterministic execution and communication rates. As machines and networks exhibit nondeterministic behavior, this static model exhibits schedule times that deviate from the deterministically predicted values. The nondeterministic model enables an additional prediction of schedule behavior based on network variance. This nondeterministic prediction better describes the average observed schedule length than the deterministic prediction. The ability of the statistical model to accurately depict observed variance is currently

under investigation. If It is hoped that future extensions of the scheduling model will incorporate the nondeterministic model into the algorithm decomposition method.

Acknowledgements

The authors would like to thank the reviewers for their time and helpful suggestions and also the University of Cincinnati Space Engineering Research Center for support of this project.

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